

$$\hat{y} = \alpha + \beta x_i + u_i$$

Assumptions for random term (u) for applying OLS method :-

1. u is purely a random term value of u depends upon in chances.

2. The mean value of u is always 0

$$\text{i.e. mean of } u = \sum (u_i) = 0$$

u can take + & - value so that their sum is 0.

3. The variance of u for different values of x always remain constant i.e. -

$$\text{variance of } u_i = E(u_i - E(u_i))^2 = \text{Constant let } \sigma_u^2$$

4. The values of u for any particular value of x is normally distributed i.e. $u_i \sim N(0, \sigma_u^2)$.

5. The different values of u are not correlated with each other i.e.

$$E(u_i, u_j) = 0$$

$$\text{Covariance}(u_i, u_j) = 0$$

6. The random term u is not correlated with the explanatory variable x i.e. $\text{cov}(u_i, x_i) = 0$ & $E(u_i, x_i) = 0$

Other assumptions

- ② The values of n remain the same in repeated samples.
- ③ Different explanatory variables are not correlated with each other i.e. $E(u_i; u_j) = 0$

Other assumptions

- ② The data is measured correctly.
- ③ Model is correctly specified.
- ④ The model is identified.

OLS Method

The method of OLS is attributed to Carl Friedrich Gauss, a German Mathematician.

The two variable PRF

$$Y_i = \alpha + \beta X_i + u_i$$

The PRF is not directly observable. we estimate it from SRF

$$Y_i = \hat{\alpha} + \hat{\beta} X_i + \hat{u}_i$$

$$Y_i = \hat{y}_i + \hat{u}_i \quad \left[\hat{y}_i \text{ is the estimated value} \right]$$

Now

$$\begin{aligned} \hat{u}_i &= Y_i - \hat{y}_i \\ &= Y_i - \hat{\alpha} - \hat{\beta} X_i \end{aligned}$$

which shows that \hat{u}_i (the residuals) are simply the difference between actual and estimated Y values.

Now given n pairs of observations on Y and X , we would like to determine SRF in such a manner that it is as close as possible to the actual Y . i.e. the sum of residuals $\sum \hat{u}_i = \sum (Y_i - \hat{Y}_i)$ is as small as possible.

It is not good criterion if we adopt the criterion of minimising $\sum \hat{u}_i$ because the algebraic sum of these residuals is zero (i.e. $\sum \hat{u}_i = 0$). We can avoid this ~~the~~ problem if we adopt the least-squares criterion.

$$\begin{aligned} \sum \hat{u}_i^2 &= \sum (Y_i - \hat{Y}_i)^2 \\ &= \sum (Y_i - \hat{\alpha} - \hat{\beta} X_i)^2 \end{aligned}$$

Now $\sum \hat{u}_i^2 = f(\hat{\alpha}, \hat{\beta})$

The method of least squares provides us with unique estimates of $\hat{\alpha}$ and $\hat{\beta}$ that give the smallest possible value of $\sum \hat{u}_i^2$.

Differentiating $\sum \hat{u}_i^2$ by α and β we get —

$$\frac{d \sum \hat{u}_i^2}{d \hat{\alpha}} = \frac{d \sum (Y_i - \hat{\alpha} - \hat{\beta} X_i)^2}{d \hat{\alpha}} = 0$$

$$\Rightarrow -2 \sum (Y_i - \hat{\alpha} - \hat{\beta} X_i) = 0$$

$$\Rightarrow \sum Y_i - n \hat{\alpha} - \hat{\beta} \sum X_i = 0 \quad \text{--- (i)}$$

⇒

$$(ii) \frac{d \sum u_i^2}{d \hat{\beta}} = \frac{d}{d \hat{\beta}} \sum (y_i - \hat{\alpha} - \hat{\beta} x_i)^2 = 0$$

$$\Rightarrow -2 \sum (y_i - \hat{\alpha} - \hat{\beta} x_i) \cdot -x_i = 0$$

$$\Rightarrow \sum x_i \sum (y_i - \hat{\alpha} - \hat{\beta} x_i) = 0$$

$$\Rightarrow \sum x_i y_i - \hat{\alpha} \sum x_i - \hat{\beta} \sum x_i^2 = 0 \quad \text{--- (ii)}$$

Now, dividing (i) by n , we get

$$\Rightarrow \frac{\sum y_i}{n} - \frac{\sum x_i \hat{\alpha}}{\sum x_i} - \frac{\hat{\beta} \sum x_i^2}{\sum x_i} = 0$$

$$\Rightarrow \bar{y} - \hat{\alpha} - \hat{\beta} \bar{x} = 0$$

$$\Rightarrow \hat{\alpha} = \bar{y} - \hat{\beta} \bar{x}$$

Now, putting $\hat{\alpha} = \bar{y} - \hat{\beta} \bar{x}$, in (ii) we get

$$\Rightarrow \sum x_i y_i - (\bar{y} - \hat{\beta} \bar{x}) \sum x_i - \hat{\beta} \sum x_i^2 = 0$$

$$\Rightarrow \sum x_i y_i - n \bar{y} \bar{x} + n \hat{\beta} \bar{x}^2 - \hat{\beta} \sum x_i^2 = 0$$

$$\Rightarrow \sum x_i y_i - (n \bar{y} \bar{x} - \hat{\beta} (\sum x_i^2 - n \bar{x}^2)) = 0$$

$$\Rightarrow \hat{\beta} = \frac{\sum x_i y_i - n \bar{y} \bar{x}}{\sum x_i^2 - n \bar{x}^2}$$

$$\Rightarrow \hat{\beta} = \frac{\sum x_i y_i - n \frac{\sum y_i}{n} \frac{\sum x_i}{n}}{\sum x_i^2 - n \left(\frac{\sum x_i}{n}\right)^2}$$

$$\Rightarrow \hat{\beta} = \frac{n \sum x_i y_i - \sum x_i \sum y_i}{n \sum x_i^2 - (\sum x_i)^2}$$

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$$\Rightarrow \hat{\beta} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2}$$

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Now, the properties of the weights k_i are given that $\hat{\beta}$ is a linear estimator because $\hat{\beta}$ is a linear function of y .
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$$\sum k_i = 1 \quad \sum k_i x_i = \bar{x}$$